

Performance analysis in wireless communications and large deviations of extreme eigenvalues of deformed random matrices

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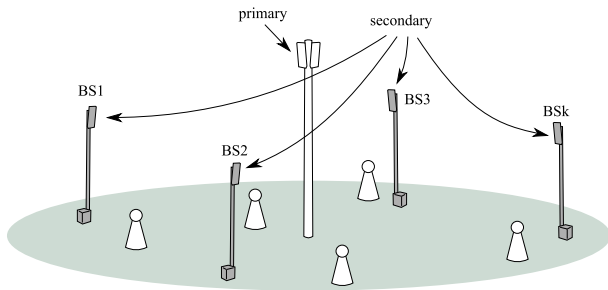
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 - ▶ General results
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 - ▶ Application to some classical models
- ▶ Conclusion

Source detection in cooperative spectrum sensing



Secondary sensors try to find a bandwidth to occupy. Those **K** sensors can share information, each of them receiving **N** samples of the signal.

Modelisation of the statistical test

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We want to test

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$$y_k(n) = w_k(n) , \quad n = 1 \dots N$$

where $w_k(n) \sim \mathcal{CN}(0, \sigma^2)$ is a white noise.

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- **Hypothesis H1 : Presence of a signal.** The data received by sensor number k is now of the form :

$$y_k(n) = h_k s(n) + w_k(n) , \quad n = 1 \dots N$$

where $s(n)$ is a Gaussian primary signal and h_k the fading coefficient associated to the secondary sensor k .

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As σ and \mathbf{h} are unknown, the Neyman-Pearson test cannot be implemented.

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$$p_0(\mathbf{Y}; \sigma^2) = (\pi\sigma^2)^{-NK} \exp\left(-\frac{N}{\sigma^2} \text{tr } \mathbf{R}\right) .$$

where $\mathbf{R} = \frac{1}{N} \mathbf{Y} \mathbf{Y}^*$ is the empirical covariance matrix.

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- Under \mathbf{H}_1 , the column vectors of \mathbf{Y} are i.i.d. $\mathcal{CN}(0, \mathbf{h} \mathbf{h}^* + \sigma^2 \mathbf{I}_K)$ where $\mathbf{h} = [h_1, \dots, h_K]^T$ is the fading vector corresponding to the K secondary sensors. The likelihood writes :

$$p_1(\mathbf{Y}; \mathbf{h}, \sigma^2) = (\pi^K \det(\mathbf{h} \mathbf{h}^* + \sigma^2 \mathbf{I}_K))^{-N} \exp\left(-N \text{tr}(\mathbf{R}(\mathbf{h} \mathbf{h}^* + \sigma^2 \mathbf{I}_K)^{-1})\right) .$$

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After some standard computations, we get the following test :

Reject \mathbf{H}_0 whenever the statistics :

$$T_N := \frac{\lambda_{\max}}{\frac{1}{K} \text{tr} \mathbf{R}}$$

is **above** the threshold γ

where λ_{\max} is the largest eigenvalue of $\mathbf{R} := \frac{1}{N} \mathbf{Y} \mathbf{Y}^*$.

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The Receiver Operating Characteristic (ROC curve) is the set of points (Type I Error, Type II Error) for all possible thresholds.

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\Rightarrow We study the ROC curve in the asymptotic regime :

$$K \rightarrow \infty, N \rightarrow \infty, \frac{K}{N} \rightarrow c \in (0, 1)$$

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- We get that, if $T_N = \frac{\lambda_{\max}}{\frac{1}{K} \text{tr} \mathbf{R}}$ and $c_N = \frac{K}{N}$,

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⇒ This determines the asymptotic threshold γ for a fixed PFA.

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- ▶ Consequently, $T_N = \frac{\lambda_{\max}}{\frac{1}{K} \text{tr} \mathbf{R}}$ converges to

$$\lambda_{\text{spiked}} := (1 + \rho) \left(1 + \frac{c}{\rho} \right) > (1 + \sqrt{c})^2 := \lambda^+ .$$

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Otherwise stated,

$$\begin{aligned} P_0[T_N > \gamma] &\simeq e^{-N \mathcal{E}_0(\gamma)} \\ P_1[T_N < \gamma] &\simeq e^{-N \mathcal{E}_1(\gamma)} . \end{aligned}$$

The set of couples $(\mathcal{E}_0(\gamma), \mathcal{E}_1(\gamma))$ is called **asymptotic error exponent curve**

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- Deviations of λ_{\max} under H_0 (cf Ben Arous, Dembo, Guionnet)
Deviations of λ_{\max} under H_1 (“spiked” model) (cf Maïda)

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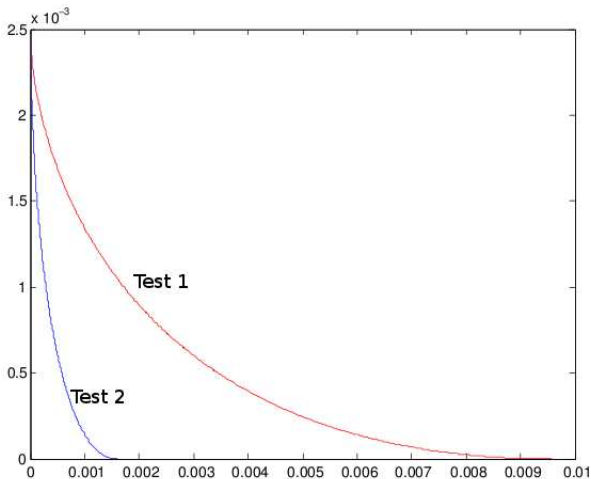
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This means that for any open set $O \subset \mathbb{R}^{r_0}$,

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Important generalisation : we can relax the hypothesis on the extreme eigenvalues, provided the law of $\frac{G}{\sqrt{n}}$ satisfies a LDP.

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Consider the i.i.d. case when $X_n = 0$. If G_n are $n \times r$ matrices whose rows are i.i.d. copies of G and $\Theta = \text{diag}(\theta_1, \dots, \theta_r)$, we can study the eigenvalues of $W_n = \frac{1}{n} G_n^* \Theta G_n$ (see Fey, van der Hofstad, Klok, $\Theta = Id$).

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In the standard case,

$$L(\alpha_1, \dots, \alpha_r) = \frac{1}{2} \sum_{i=1}^r \left(\frac{\alpha_i}{\theta_i} - 1 - \log \frac{\alpha_i}{\theta_i} \right).$$

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From there it is easy to deduce the rate function for the largest eigenvalue

$$L_{\max}(x) = \begin{cases} \frac{1}{2}(x - 1 - \log x) & \text{if } x \geq 1 \\ \frac{r}{2}(x - 1 - \log x) & \text{if } x \in (0, 1) \end{cases}$$

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Rq : we first condition on the deviation of the eigenvalues of X_n so that we can consider those as outliers.

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$$f_n(z) = \det \left([G_{i,j}^n(z)]_{i,j=1}^r - \text{diag}(\theta_1^{-1}, \dots, \theta_r^{-1}) \right),$$

with

$$G_{i,j}^n(z) = \langle U_i^n, (z - X_n)^{-1} U_j^n \rangle.$$

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If

$$K_{i,j}^n(z) = \langle G_i^n, (z - X_n)^{-1} G_j^n \rangle = \frac{1}{n} \sum_{k=1}^n \frac{g_i(k) \overline{g_j(k)}}{z - \lambda_k}$$

and $C_{i,j}^n = \frac{1}{n} \sum_{k=1}^n g_i(k) \overline{g_j(k)}$ then $f_n(z) = P_{\Theta,r}(K^n(z), C^n)$

Conclusion

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Can we use large deviation principles of this type to analyse the performance for some other models relevant in wireless communication context?